

# Dynamic Approaches: The Hidden Markov Model

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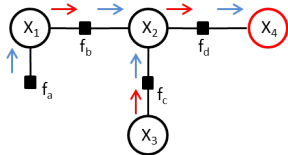
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Machine Learning: Neural Networks and Advanced Models  
(AA2)



# Inference as Message Passing

How to infer the distribution  $P(\mathbf{X}_{unk} | \mathbf{X}_{obs})$  of a number of random variables  $\mathbf{X}_{unk}$  in the graphical model, given the observed values of other variables  $\mathbf{X}_{obs}$



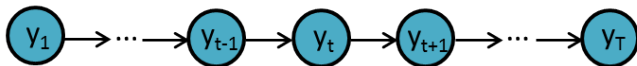
Directed and undirected models of **fixed structure**

- Exact inference
  - Passing **messages** (vectors of information) on the structure of the graphical model following a **propagation direction**
  - Works for chains, trees and can be used in (some) graphs
- Approximated inference can use approximations of the distribution (**variational**) or can estimate its expectation using examples (**sampling**)

## Today's Lecture

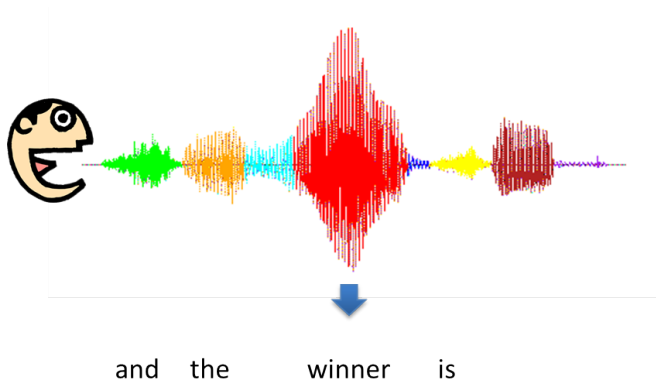
- Exact inference on a **chain** with **observed and unobserved** variables
- A probabilistic model for sequences: **Hidden Markov Models** (HMMs)
- Using inference to **learn**: the Expectation-Maximization algorithm for HMMs
- Graphical models with **varying structure**: Dynamic Bayesian Networks
- Application examples

# Sequences

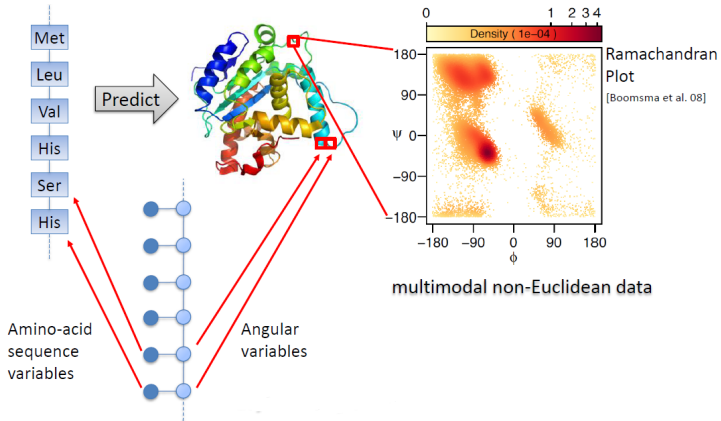


- A sequence  $\mathbf{y}$  is a collection of observations  $y_t$  where  $t$  represent the **position of the element** according to a (complete) order (e.g. **time**)
- Reference population is a set of i.i.d sequences  $\mathbf{y}^1, \dots, \mathbf{y}^N$
- Different sequences  $\mathbf{y}^1, \dots, \mathbf{y}^N$  generally have **different lengths**  $T^1, \dots, T^N$

# Sequences in Speech Processing



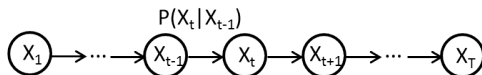
# Sequences in Biology



# Markov Chain

## First-Order Markov Chain

Directed graphical model for sequences s.t. element  $X_t$  only depends on its predecessor in the sequence



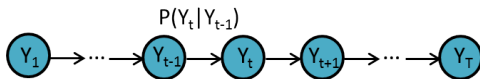
- Joint probability **factorizes** as

$$P(\mathbf{X}) = P(X_1, \dots, X_T) = P(X_1) \prod_{t=2}^T P(X_t | X_{t-1})$$

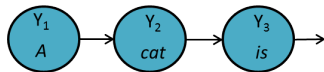
- $P(X_t | X_{t-1})$  is the **transition distribution**;  $P(X_1)$  is the **prior distribution**
- General form: an  **$L$ -th order Markov chain** is such that  $X_t$  depends on  $L$  predecessors

# Observed Markov Chains

Can we use a Markov chain to model the relationship between observed elements in a sequence?



Of course yes, but...

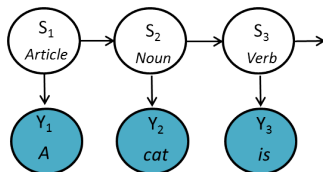


Does it make sense to represent  $P(is|cat)$ ?



# Hidden Markov Model (HMM) (I)

Stochastic process where **transition dynamics** is **disentangled from observations** generated by the process

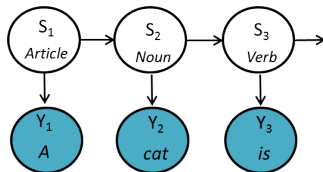


- **State transition is an unobserved** (hidden/latent) process characterized by the **hidden state variables**  $S_t$ 
  - $S_t$  are often **discrete** with value in  $\{1, \dots, C\}$
  - Multinomial **state transition** and prior probability (**stationarity assumption**)

$$A_{ij} = P(S_t = i | S_{t-1} = j) \text{ and } \pi_i = P(S_t = i)$$

# Hidden Markov Model (HMM) (II)

Stochastic process where **transition dynamics** is **disentangled from observations** generated by the process



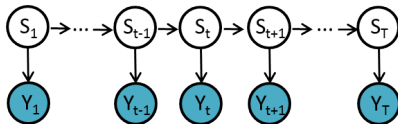
- Observations are generated by the **emission distribution**

$$b_i(y_t) = P(Y_t = y_t | S_t = i)$$

# HMM Joint Probability Factorization

Discrete-state HMMs are **parameterized** by  $\theta = (\pi, A, B)$  and the **finite number of hidden states**  $C$

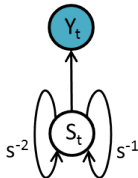
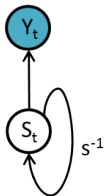
- State transition and prior distribution  $A$  and  $\pi$
- Emission distribution  $B$  (or its parameters)



$$\begin{aligned} P(\mathbf{Y} = \mathbf{y}) &= \sum_{\mathbf{s}} P(\mathbf{Y} = \mathbf{y}, \mathbf{S} = \mathbf{s}) \\ &= \sum_{s_1, \dots, s_T} \left\{ P(S_1 = s_1) P(Y_1 = y_1 | S_1 = s_1) \right. \\ &\quad \left. \prod_{t=2}^T P(S_t = s_t | S_{t-1} = s_{t-1}) P(Y_t = y_t | S_t = s_t) \right\} \end{aligned}$$

# HMMs as a Recursive Model

A graphical framework describing **how contextual information is recursively encoded** by both probabilistic and neural models



- Indicates that the hidden state  $S_t$  at time  $t$  is dependent on **context information** from
  - The previous time step  $s^{-1}$
  - Two time steps earlier  $s^{-2}$
  - ...
- When applying the recursive model to a sequence (**unfolding**), it generates the corresponding **directed graphical model**

## 3 Notable Inference Problems

### Definition (Smoothing)

Given a model  $\theta$  and an observed sequence  $\mathbf{y}$ , determine the **distribution of the  $t$ -th hidden state**  $P(S_t | \mathbf{Y} = \mathbf{y}, \theta)$

### Definition (Learning)

Given a dataset of  $N$  observed sequences  $\mathcal{D} = \{\mathbf{y}^1, \dots, \mathbf{y}^N\}$  and the number of hidden states  $C$ , **find the parameters  $\pi$ ,  $A$  and  $B$**  that maximize the probability of model  $\theta = \{\pi, A, B\}$  having generated the sequences in  $\mathcal{D}$

### Definition (Optimal State Assignment)

Given a model  $\theta$  and an observed sequence  $\mathbf{y}$ , find an **optimal state assignment**  $\mathbf{s} = s_1^*, \dots, s_T^*$  for the hidden Markov chain

# Forward-Backward Algorithm

**Smoothing** - How do we determine the posterior  $P(S_t = i | \mathbf{y})$ ?

Exploit factorization

$$\begin{aligned} P(S_t = i | \mathbf{y}) &\propto P(S_t = i, \mathbf{y}) = P(S_t = i, \mathbf{Y}_{1:t}, \mathbf{Y}_{t+1:T}) \\ &= P(S_t = i, \mathbf{Y}_{1:t}) P(\mathbf{Y}_{t+1:T} | S_t = i) = \alpha_t(i) \beta_t(i) \end{aligned}$$

$\alpha$ -term computed as part of **forward recursion** ( $\alpha_1(i) = b_i(y_1) \pi_i$ )

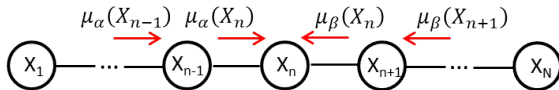
$$\alpha_t(i) = P(S_t = i, \mathbf{Y}_{1:t}) = b_i(y_t) \sum_{j=1}^C A_{ij} \alpha_{t-1}(j)$$

$\beta$ -term computed as part of **backward recursion** ( $\beta_T(i) = 1, \forall i$ )

$$\beta_t(j) = P(\mathbf{Y}_{t+1:T} | S_t = j) = \sum_{i=1}^C b_i(y_{t+1}) \beta_{t+1}(i) A_{ij}$$

# Deja vu

Doesn't the **Forward-Backward algorithm** look strangely familiar?



- $\alpha_t \equiv \mu_\alpha(X_n) \rightarrow$  forward message

$$\underbrace{\mu_\alpha(X_n)}_{\alpha_t(i)} = \sum_{\substack{X_{n-1} \\ \sum_{j=1}^C}} \underbrace{\psi(X_{n-1}, X_n)}_{b_i(y_t)A_{ij}} \underbrace{\mu_\alpha(X_{n-1})}_{\alpha_{t-1}(j)}$$

- $\beta_t \equiv \mu_\beta(X_n) \rightarrow$  backward message

$$\underbrace{\mu_\beta(X_n)}_{\beta_t(j)} = \sum_{\substack{X_{n+1} \\ \sum_{i=1}^C}} \underbrace{\psi(X_n, X_{n+1})}_{b_i(y_{t+1})A_{ij}} \underbrace{\mu_\beta(X_{n+1})}_{\beta_{t+1}(i)}$$

# Learning in HMM

Learning HMM parameters  $\theta = (\pi, \mathbf{A}, \mathbf{B})$  by **maximum likelihood**

$$\begin{aligned}\mathcal{L}(\theta) &= \log \prod_{n=1}^N P(\mathbf{Y}^n | \theta) \\ &= \log \prod_{n=1}^N \left\{ \sum_{s_1^n, \dots, s_{T_n}^n} P(S_1^n) P(Y_1^n | S_1^n) \prod_{t=2}^{T_n} P(S_t^n | S_{t-1}^n) P(Y_t^n | S_t^n) \right\}\end{aligned}$$

- How can we deal with the unobserved random variables  $S_t$  and the nasty summation in the log?
- Expectation-Maximization algorithm
  - Maximization of the **complete likelihood**  $\mathcal{L}_c(\theta)$
  - Completed with **indicator variables**

$$z_{it}^n = \begin{cases} 1 & \text{if } n\text{-th chain is in state } i \text{ at time } t \\ 0 & \text{otherwise} \end{cases}$$



# Complete HMM Likelihood

Introduce indicator variables in  $\mathcal{L}(\theta)$  together with model parameters  $\theta = (\pi, \mathbf{A}, \mathbf{B})$

$$\begin{aligned}\mathcal{L}_c(\theta) &= \log P(\mathcal{X}, \mathcal{Z}|\theta) = \log \prod_{n=1}^N \left\{ \prod_{i=1}^C [P(S_1^n = i)P(Y_1^n|S_1^n = i)]^{z_{1i}^n} \right. \\ &\quad \left. \prod_{t=2}^{T_n} \prod_{i,j=1}^C P(S_t^n = i|S_{t-1}^n = j)^{z_{ij}^n z_{(t-1)j}^n} P(Y_t^n|S_t^n = i)^{z_{ti}^n} \right\} \\ &= \sum_{n=1}^N \left\{ \sum_{i=1}^C z_{1i}^n \log \pi_i + \sum_{t=2}^{T_n} \sum_{i,j=1}^C z_{ij}^n z_{(t-1)j}^n \log A_{ij} + \sum_{t=1}^{T_n} \sum_{i=1}^C z_{ti}^n \log b_i(y_t^n) \right\}\end{aligned}$$

# Expectation-Maximization

A 2-step iterative algorithm for the maximization of **complete likelihood**  $\mathcal{L}_c(\theta)$  w.r.t. model parameters  $\theta$

**E-Step:** Given the current estimate of the model parameters  $\theta^{(t)}$ , compute

$$Q^{(t+1)}(\theta|\theta^{(t)}) = E_{\mathcal{Z}|\mathcal{X},\theta^{(t)}}[\log P(\mathcal{X}, \mathcal{Z}|\theta)]$$

**M-Step:** Find the new estimate of the model parameters

$$\theta^{(t+1)} = \arg \max_{\theta} Q^{(t+1)}(\theta|\theta^{(t)})$$

Iterate 2 steps until  $|\mathcal{L}_c(\theta)^{it} - \mathcal{L}_c(\theta)^{it-1}| < \epsilon$  (or stop if maximum number of iterations is reached)

# E-Step (I)

Compute the expected value expectation of the complete log-likelihood w.r.t indicator variables  $z_i^n$  assuming (estimated) parameters  $\theta^t = (\pi^t, A^t, B^t)$  fixed at time  $t$  (i.e. constants)

$$Q^{(t+1)}(\theta|\theta^{(t)}) = E_{\mathcal{Z}|\mathcal{X},\theta^{(t)}}[\log P(\mathcal{X}, \mathcal{Z}|\theta)]$$

Expectation w.r.t a (discrete) random variable  $z$  is

$$E_Z[Z] = \sum_z z \cdot P(Z = z)$$

To compute the conditional expectation  $Q^{(t+1)}(\theta|\theta^{(t)})$  for the complete HMM log-likelihood we need to estimate

$$E_{\mathcal{Z}|\mathbf{Y},\theta^{(k)}}[z_{ti}] = P(S_t = i|\mathbf{y})$$

$$E_{\mathcal{Z}|\mathbf{Y},\theta^{(k)}}[z_{ti}z_{(t-1)j}] = P(S_t = i, S_{t-1} = j|\mathbf{Y})$$

## E-Step (II)

We know how to compute the posteriors by the **forward-backward algorithm!**

$$\gamma_t(i) = P(S_t = i | \mathbf{Y}) = \frac{\alpha_t(i)\beta_t(i)}{\sum_{j=1}^C \alpha_t(j)\beta_t(j)}$$

$$\gamma_{t,t-1}(i,j) = P(S_t = i, S_{t-1} = j | \mathbf{Y}) = \frac{\alpha_{t-1}(j)A_{ij}b_i(y_t)\beta_t(i)}{\sum_{m,l=1}^C \alpha_{t-1}(m)A_{lm}b_l(y_t)\beta_t(l)}$$

# M-Step (I)

Solve the **optimization problem**

$$\theta^{(t+1)} = \arg \max_{\theta} Q^{(t+1)}(\theta | \theta^{(t)})$$

using the information computed at the E-Step (the posteriors).

**How?**

As usual

$$\frac{\partial Q^{(t+1)}(\theta | \theta^{(t)})}{\partial \theta}$$

where  $\theta = (\pi, A, B)$  are now variables.

## Attention

Parameters can be distributions  $\Rightarrow$  need to preserve sum-to-one constraints (**Lagrange Multipliers**)

# M-Step (II)

State distributions

$$A_{ij} = \frac{\sum_{n=1}^N \sum_{t=2}^{T^n} \gamma_{t,t-1}^n(i,j)}{\sum_{n=1}^N \sum_{t=2}^{T^n} \gamma_{t-1}^n(j)} \quad \text{and} \quad \pi_i = \sum_{n=1}^N \gamma_1^n(i)$$

Emission distribution (multinomial)

$$B_{ki} = \sum_{n=1}^N \sum_{t=1}^{T_n} \gamma_t^n(i) \delta(y_t = k)$$

where  $\delta(\cdot)$  is the indicator function for emission symbols  $k$

# Decoding Problem

- Find the **optimal hidden state assignment**  $\mathbf{s} = s_1^*, \dots, s_T^*$  for an observed sequence  $\mathbf{y}$  given a trained HMM
- No unique interpretation of the problem

- Identify the **single hidden states**  $s_t$  that maximize the posterior

$$s_t^* = \arg \max_{i=1, \dots, C} P(S_t = i | \mathbf{Y})$$

- Find the most likely **joint hidden state assignment**

$$\mathbf{s}^* = \arg \max_{\mathbf{s}} P(\mathbf{Y}, \mathbf{S} = \mathbf{s})$$

- The last problem is addressed by the **Viterbi algorithm**

# Viterbi Algorithm

An efficient **dynamic programming** algorithm based on a **backward-forward recursion**

An example of a **max-product message passing** algorithm

Recursive backward term

$$\epsilon_{t-1}(s_{t-1}) = \max_{s_t} P(Y_t | S_t = s_t) P(S_t = s_t | S_{t-1} = s_{t-1}) \epsilon_t(s_t),$$

Root optimal state

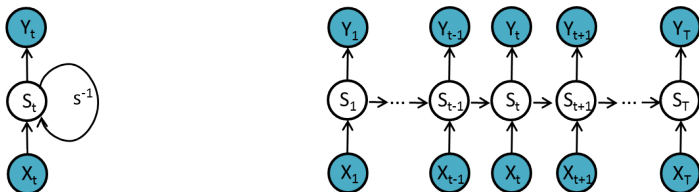
$$s_1^* = \arg \max_s P(Y_t | S_1 = s) P(S_1 = s_1) \epsilon_1(s).$$

Recursive forward optimal state

$$s_t^* = \arg \max_s P(Y_t | S_t = s) P(S_t = s | S_{t-1} = s_{t-1}^*) \epsilon_t(s).$$



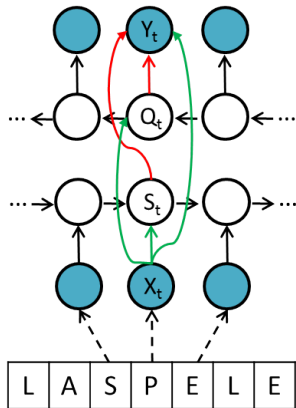
# Input-Output Hidden Markov Models



- Translate an input sequence into an output sequence (**transduction**)
- State transition and emissions depend on input observations (**input-driven**)
- Recursive model highlights analogy with **recurrent neural networks**

# Bidirectional Input-driven Models

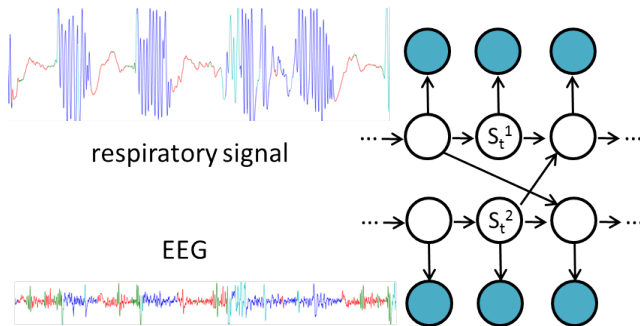
Remove **causality assumption** that current observation does not depend on the future and **homogeneity assumption** that a state transition is not dependent on the position in the sequence



- Structure and function of a region of DNA and protein sequences may depend on upstream and downstream information
- Hidden state transition distribution changes with the amino-acid sequence being fed in input

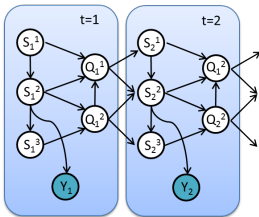
# Coupled HMM

Describing **interacting processes** whose observations follow different dynamics while the underlying generative processes are interlaced

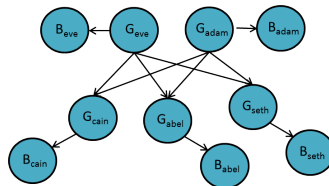


# Dynamic Bayesian Networks

HMMs are a specific case of a class of directed models that represent **dynamic processes** and data with **changing connectivity template**



Hierarchical HMM



Structure changing information

## Dynamic Bayesian Networks (DBN)

Graphical models whose structure changes to represent evolution across time and/or between different samples

# Take Home Messages

- Hidden Markov Models
  - Hidden states used to realize an **unobserved generative process for sequential data**
  - A **mixture model** where selection of the next component is regulated by the transition distribution
  - Hidden states **summarize (cluster) information** on subsequences in the data
- Inference in HMMS
  - **Forward-backward** - Hidden state posterior estimation
  - **Expectation-Maximization** - HMM parameter learning
  - **Viterbi** - Most likely hidden state sequence
- Dynamic Bayesian Networks
  - A graphical **model whose structure changes** to reflect information with variable size and connectivity patterns
  - Suitable for modeling **structured data** (sequences, tree, ...)