

Master Program in *Data Science and Business Informatics*

Statistics for Data Science

Lesson 27 - Bootstrap and resampling methods

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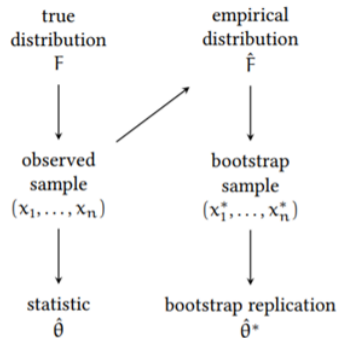
Bootstrap principle

- Let $X_1, \dots, X_n \sim F$ be a random sample
 - ▶ with *unknown distribution* F
- Estimator $T = h(X_1, \dots, X_n)$, e.g., $\bar{X}_n = (X_1 + \dots + X_n)/n$
 - ▶ with *unknown (sampling) distribution*
- From a dataset x_1, \dots, x_n , we can derive a point estimate $\hat{\theta} = h(x_1, \dots, x_n)$
- From **many datasets** $\{x_1^i, \dots, x_n^i\}_{i=1}^m$, we can derive many point estimates $\hat{\theta}^i = h(x_1^i, \dots, x_n^i)$
- By the **Glivenko-Cantelli Thm**, the empirical distribution of $\hat{\theta}^i$ approximates the distribution of T
- **Problem:** typically, we do not have many datasets, but only one!

See R script

Bootstrap principle

- Let $X_1, \dots, X_n \sim F$ be a random sample
 - ▶ with *unknown distribution* F
 - Estimator $T = h(X_1, \dots, X_n)$, e.g., $\bar{X}_n = (X_1 + \dots + X_n)/n$
 - From a dataset x_1, \dots, x_n , we can
 - ▶ derive a point estimate $\hat{\theta} = h(x_1, \dots, x_n)$
 - ▶ or, **derive an estimate \hat{F} of F**
 - From \hat{F} we can generate (a lot of) *bootstrap samples* x_1^*, \dots, x_n^*
 - ▶ as realizations of $X_1^*, \dots, X_n^* \sim \hat{F}$
- and then (many) bootstrap point estimates $\hat{\theta}^* = h(x_1^*, \dots, x_n^*)$



BOOTSTRAP PRINCIPLE. Use the dataset x_1, x_2, \dots, x_n to compute an estimate \hat{F} for the “true” distribution function F . Replace the random sample X_1, X_2, \dots, X_n from F by a random sample $X_1^*, X_2^*, \dots, X_n^*$ from \hat{F} , and approximate the probability distribution of $h(X_1, X_2, \dots, X_n)$ by that of $h(X_1^*, X_2^*, \dots, X_n^*)$.

Empirical bootstrap

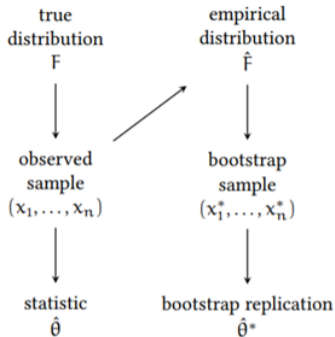
- How to derive \hat{F} from x_1, \dots, x_n ?
- If we know nothing about F , use the empirical distribution:

$$\hat{F}(a) = F_n(a) = \frac{|\{i \in \{1, \dots, n\} \mid x_i \leq a\}|}{n}$$

- How to generate a bootstrap sample x_1^*, \dots, x_n^* ?
 - ▶ x_i^* is chosen randomly from \hat{F}
 - ▶ i.e., x_i^* s chosen randomly from x_1, \dots, x_n (our dataset)
- Hence, a bootstrap dataset x_1^*, \dots, x_n^* is obtained by *random sampling with replacement!*
- Often the bootstrap approximation of the distribution of T will improve if we shift T by relating it to a corresponding feature of the “true” distribution.

- ▶ rather than approximating the distribution of \bar{X}_n by the one of \bar{X}_n^* , better to approximate $\Delta = \bar{X}_n - \mu$ by $\Delta^* = \bar{X}_n^* - \mu^*$, where $\mu^* = E[\hat{F}] = \bar{x}_n = (x_1 + \dots + x_n)/n$

[See remarks 18.1 and 18.2 of textbook]



Empirical bootstrap

EMPIRICAL BOOTSTRAP SIMULATION (FOR $\bar{X}_n - \mu$). Given a dataset x_1, x_2, \dots, x_n , determine its empirical distribution function F_n as an estimate of F , and compute the expectation

$$\mu^* = \bar{x}_n = \frac{x_1 + x_2 + \dots + x_n}{n}$$

corresponding to F_n .

1. Generate a bootstrap dataset $x_1^*, x_2^*, \dots, x_n^*$ from F_n .
2. Compute the centered sample mean for the bootstrap dataset:

$$\bar{x}_n^* - \bar{x}_n,$$

where

$$\bar{x}_n^* = \frac{x_1^* + x_2^* + \dots + x_n^*}{n}.$$

Repeat steps 1 and 2 many times.

- Use the empirical distribution of $\delta^* = \bar{x}_n^* - \bar{x}_n$ (realizations of $\Delta^* = \bar{X}_n^* - \bar{x}_n$)
 - ▶ for estimating the distribution of $\Delta = \bar{X}_n - \mu$, and in particular:

$$E[\Delta] = E[\bar{X}_n] - \mu \approx E[\Delta^*] \approx \text{mean}(\delta^*)$$

- ▶ and then estimate μ as $\hat{\mu} = E[\bar{X}_n] - \text{mean}(\delta^*) \approx \bar{x}_n - \text{mean}(\delta^*)$

$\text{mean}(\delta^*)$ is the estimated bias

- ▶ and $\text{se}(\bar{X}_n) = \sqrt{\text{Var}(\bar{X}_n)} = \sqrt{\text{Var}(\bar{X}_n - \mu)} \approx \sqrt{\text{Var}(\bar{X}_n^* - \bar{x}_n)} \approx \text{sd}(\delta^*)$

See R script

Empirical bootstrap

EMPIRICAL BOOTSTRAP SIMULATION (FOR $\bar{X}_n - \mu$). Given a dataset x_1, x_2, \dots, x_n , determine its empirical distribution function F_n as an estimate of F , and compute the expectation

$$\mu^* = \bar{x}_n = \frac{x_1 + x_2 + \dots + x_n}{n}$$

corresponding to F_n .

1. Generate a bootstrap dataset $x_1^*, x_2^*, \dots, x_n^*$ from F_n .
2. Compute the centered sample mean for the bootstrap dataset:

$$\bar{x}_n^* - \bar{x}_n,$$

where

$$\bar{x}_n^* = \frac{x_1^* + x_2^* + \dots + x_n^*}{n}.$$

Repeat steps 1 and 2 many times.

- Use the empirical distribution of $\delta^* = \bar{x}_n^* - \bar{x}_n$ (realizations of $\Delta^* = \bar{X}_n^* - \bar{x}_n$)
 - ▶ for estimating the distribution of $\Delta = \bar{X}_n - \mu$, and in particular:
 - ▶ confidence interval for $\delta = \bar{x}_n - \mu$ is $(q_{\alpha/2}, q_{1-\alpha/2})$ of δ^* empirical distribution
 - ▶ $q_{\alpha/2} \leq \delta = \bar{x}_n - \mu \leq q_{1-\alpha/2}$ implies c.i. for μ is $(\bar{x}_n - q_{1-\alpha/2}, \bar{x}_n - q_{\alpha/2})$

See R script

Empirical bootstrap

`boot.ci` method in R confidence intervals:

- `type='basic'`: $(\bar{x}_n - q_{1-\alpha/2}, \bar{x}_n - q_{\alpha/2})$ with quantiles over the distribution of δ^*
- `type='perc'`: $(q_{\alpha/2}, q_{1-\alpha/2})$ with quantiles over the distribution of \bar{x}_n^* (without shift)
- `type='norm'`: $(\bar{x}_n - q_{1-\alpha/2}, \bar{x}_n - q_{\alpha/2})$ with quantiles over $N(\text{mean}(\delta^*), \text{var}(\delta^*))$
- `type='bca'`: bias (and skewness) correction and acceleration

See R script

Empirical bootstrap

`boot.ci` method in R confidence intervals:

- `type='stud'`: $(\bar{x}_n - q_{1-\alpha/2} \frac{s_n}{\sqrt{n}}, \bar{x}_n - q_{\alpha/2} \frac{s_n}{\sqrt{n}})$ with quantiles over the distribution of t^*

EMPIRICAL BOOTSTRAP SIMULATION FOR THE STUDENTIZED MEAN.

Given a dataset x_1, x_2, \dots, x_n , determine its empirical distribution function F_n as an estimate of F . The expectation corresponding to F_n is $\mu^* = \bar{x}_n$.

1. Generate a bootstrap dataset $x_1^*, x_2^*, \dots, x_n^*$ from F_n .
2. Compute the studentized mean for the bootstrap dataset:

$$t^* = \frac{\bar{x}_n^* - \bar{x}_n}{s_n^*/\sqrt{n}},$$

where \bar{x}_n^* and s_n^* are the sample mean and sample standard deviation of $x_1^*, x_2^*, \dots, x_n^*$.

Repeat steps 1 and 2 many times.

See R script

Empirical bootstrap

- Bootstrap approach applies to **any** estimator, not only the mean
- Example 1: the German Tank problem

[see Lesson 19]

$$T_2 = \frac{n+1}{n} M_n - 1 \qquad E[T_2] = N$$

- Example 2: linear regression coefficients

[see Lesson 26]

- ▶ 95% confidence intervals (assuming $U_i \sim \mathcal{N}(0, \sigma^2)$):

$$\hat{\beta} \pm t_{n-2, 0.025} \text{se}(\hat{\beta}) \qquad \hat{\alpha} \pm t_{n-2, 0.025} \text{se}(\hat{\alpha})$$

See R script

An application: probability of large errors

- Bootstrap principle: for $X \sim F$
 - ▶ the empirical distribution of $\Delta^* = \bar{X}_n^* - \bar{x}_n$ approximates the distribution of $\Delta = \bar{X}_n - \mu$
- Application: estimate $P_F(|\bar{X}_n - \mu| > 1)$ as
 - ▶ $P_{\hat{F}}(|\bar{X}_n^* - \bar{x}_n| > 1)$ and then by the fraction of $\delta^* = \bar{x}_n^* - \bar{x}_n$ such that $|\delta^*| > 1$

See R script

Wrap up on empirical bootstrap

- How many bootstrap samples?
 - ▶ There are $\binom{2n-1}{n-1}$ distinct bootstrap samples *[Why?]*
 - ▶ Suggested to use at least 1000 bootstrap samples
 - ▶ **Jackknife resampling**: bootstrap samples $x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n$, for $i = 1, \dots, n$
- How good is the approximation by bootstrap?
 - ▶ Small perturbation to data-generating process should produce small perturbation of the parameter to estimate (θ)
 - ▶ Problems with extreme values, e.g., percentiles, maximum, etc.

Resampling methods for classifier performance estimation

- Decision rule $y_{\theta}^{+}(w)$ (classifier) or score function $s_{\theta}(w)$ (binary probabilistic classifier, Lesson 23)
- Loss function, e.g., 0-1 loss $\ell_{\theta}(c, w) = \mathbb{1}_{y_{\theta}^{+}(w) \neq c}$

Risk (or Expected Prediction Error EPE)

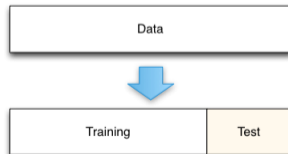
The risk w.r.t. a loss function ℓ_{θ} is $R(\theta_{TRUE}, \theta) = E_{(W, C) \sim f_{\theta_{TRUE}}} [\ell_{\theta}(C, W)]$.

Question: how to estimate risk given a dataset?

- **Holdout method:** split dataset into training and test, build $y_{\theta}^{+}()$ on training, estimate as the empirical risk on test set $(w_1, c_1), \dots, (w_n, c_n)$:

$$\hat{r} = \frac{1}{n} \sum_{i=1}^n \ell_{\theta}(c_i, w_i) \quad se = \sqrt{\frac{\hat{r}(1-\hat{r})}{n}}$$

[see Lesson 26 on CI for proportions]



- ▶ Drawbacks: variability of training/test set, and then of empirical risk estimates

Resampling methods for classifier performance estimation

Question: how to estimate risk given a dataset?

- **Random sampling:** repeat holdout k times, and average the empirical risks:

$$\hat{r} = \frac{1}{k} \sum_{j=1}^k \hat{r}^j \text{ with } \hat{r}^j = \frac{1}{n_j} \sum_{i=1}^{n_j} \ell_{\theta}(c_i^j, w_i^j) \text{ is the error on } j^{\text{th}} \text{ training-test split}$$

- Standard error calculated as standard deviation over the k repetitions:

$$se = \sqrt{\frac{1}{k-1} \sum_{j=1}^k (\hat{r}^j - \hat{r})^2}$$

Wrong! As test sets (and then \hat{r}^j 's) are not independent!

Resampling methods for classifier performance estimation

Question: how to estimate risk given a dataset?

- **k-fold cross-validation:** average the empirical risks over k -fold splits:

$$\hat{r} = \frac{1}{k} \sum_{j=1}^k \hat{r}^j \text{ with } \hat{r}^j = \frac{1}{n/k} \sum_{i=1}^{n/k} \ell_{\theta}(c_i^j, w_i^j)$$

- Standard deviation calculated over the k folds, with

$$se = \sqrt{\frac{1}{k-1} \sum_j (\hat{r}^j - \hat{r})^2}$$

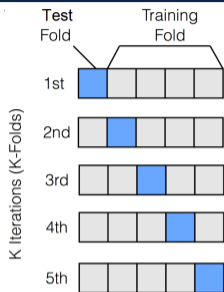
Wrong! (*) Test sets are independent, but training sets (and then \hat{r}^j 's) are not!

- If classifier is stable over the folds (see [[Kohavi, 1995](#)]), use:

$$se = \sqrt{\frac{\hat{r}(1-\hat{r})}{n}}$$

[see Lesson 26 on CI for proportions]

- ▶ Boils down to estimation as holdout but using all data instances (lower variability)!
- ▶ This is the one implemented in R/caret
- Setting $k = n$ is the **leave-one out cross-validation** (LOOCV)



(*) CV should be treated as an estimator of the average prediction error across training sets!

Resampling methods for classifier performance estimation

Question: how to estimate risk given a dataset?

- training = bootstrap x_1^*, \dots, x_n^* , test = dataset \setminus bootstrap = $\{x_1, \dots, x_n\} \setminus \{x_1^*, \dots, x_n^*\}$
 - ▶ .632 **bootstrap algorithm** for k bootstrap runs

$$\hat{r} = \frac{1}{k} \sum_j (0.632 \cdot \hat{r}^j + 0.368 \cdot \hat{r}_{tr})$$

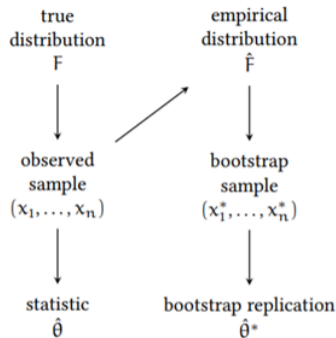
where \hat{r}^j is the empirical risk on j^{th} bootstrap run, and \hat{r}_{tr} is the empirical risk on the dataset

- [Kohavi, 1995, Kim, 2009] conclusions and recommendations:
 - ▶ Bootstrap has low variance, but it is extremely biased
 - ▶ k -fold cross-validation has low bias and variance can be controlled
 - by averaging multiple k -fold cross-validation
 - ▶ Recommendation: use **repeated (stratified) k -fold cross-validation**, with $k \approx 10$
- [Vanwinckelen, 2012] warns against “repeated”, and it recommends **k -fold cross-validation**

See R script

Parametric bootstrap principle

- Let $X_1, \dots, X_n \sim F(\gamma)$ be a random sample
 - ▶ with known family F but *unknown* parameter γ
 - Estimator $T = h(X_1, \dots, X_n)$, e.g., $\bar{X}_n = (X_1 + \dots + X_n)/n$
 - From a dataset x_1, \dots, x_n , we can
 - ▶ derive an estimate $\hat{\gamma}$ of γ
 - From $F(\hat{\gamma})$ we can generate (a lot of) *bootstrap samples* x_1^*, \dots, x_n^*
 - ▶ as realizations of $X_1^*, \dots, X_n^* \sim F(\hat{\gamma})$ [a form of **Monte Carlo simulation**]
- and then (many) bootstrap point estimates $\hat{\theta}^* = h(x_1^*, \dots, x_n^*)$



Parametric bootstrap

PARAMETRIC BOOTSTRAP SIMULATION (FOR $\bar{X}_n - \mu$). Given a dataset x_1, x_2, \dots, x_n , compute an estimate $\hat{\theta}$ for θ . Determine $F_{\hat{\theta}}$ as an estimate for F_{θ} , and compute the expectation $\mu^* = \mu_{\hat{\theta}}$ corresponding to $F_{\hat{\theta}}$.

1. Generate a bootstrap dataset $x_1^*, x_2^*, \dots, x_n^*$ from $F_{\hat{\theta}}$.
2. Compute the centered sample mean for the bootstrap dataset:

$$\bar{x}_n^* - \mu_{\hat{\theta}},$$

where

$$\bar{x}_n^* = \frac{x_1^* + x_2^* + \dots + x_n^*}{n}.$$

Repeat steps 1 and 2 many times.

- Cfr with non-parametric bootstrap: use $\mu_{\hat{\theta}}$ instead of \bar{x}_n
- Use the empirical distribution of $\delta^* = \bar{x}_n^* - \mu_{\hat{\theta}}$ for estimating
 - ▶ confidence interval for $\delta = \bar{x}_n - \mu$ is $(q_{\alpha/2}, q_{1-\alpha/2})$ of δ^* empirical distribution
 - ▶ $q_{\alpha/2} \leq \delta = \bar{x}_n - \mu \leq q_{1-\alpha/2}$ implies c.i. for μ is $(\bar{x}_n - q_{1-\alpha/2}, \bar{x}_n - q_{\alpha/2})$

See R script

Application: distribution fitting

- Consider x_1, \dots, x_n realizations of a random sample $X_1, \dots, X_n \sim F$
- Is the dataset from an $Exp(\lambda)$ for some λ ? I.e., is it $F = Exp(\lambda)$?
- We estimate $\hat{\lambda} = 1/\bar{x}_n$
- We measure how close is the dataset to the distribution as:

[MLE estimation]

$$t_{ks} = \sup_{a \in \mathbb{R}} |F_n(a) - F_{\hat{\lambda}}(a)|$$

where:

- ▶ $F_n(a)$ is the empirical cumulative distribution function of x_1, \dots, x_n
 - ▶ $F_{\hat{\lambda}}(a) = 1 - e^{-\hat{\lambda}a}$, for $a \geq 0$, is the CDF of $Exp(\hat{\lambda})$
 - ▶ t_{ks} is the *Kolmogorov-Smirnov* distance
- if $F = Exp(\lambda)$ then both $F_n \approx F$ and $F_{\hat{\lambda}} \approx F$, and then $F_n \approx F_{\hat{\lambda}}$, so that t_{ks} is small
 - if $F \neq Exp(\lambda)$ then $F_n \approx F \neq Exp(\lambda) \approx F_{\hat{\lambda}}$, so that t_{ks} is large

[See Lesson 11]

Application: distribution fitting

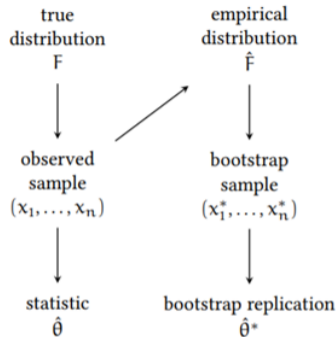
- For the software dataset from the textbook
 - ▶ $\hat{\lambda} = 0.0015$ and $t_{ks} = 0.17$
- Is $t_{ks} = 0.17$ expected or an extreme value?
- Let's study the distribution of the bootstrap estimator:

$$T_{ks} = \sup_{a \in \mathbb{R}} |F_n^*(a) - F_{\hat{\lambda}^*}(a)|$$

where:

- ▶ $X_1^*, \dots, X_n^* \sim \text{Exp}(\hat{\lambda})$ is a bootstrap sample
 - ▶ $F_n^*(a)$ is the empirical cumulative distribution of the bootstrap sample
 - ▶ $\hat{\lambda}^* = 1/\bar{X}_n^*$
- It turns out $P(T_{ks} > 0.17) \approx 0$, unlikely that $\text{Exp}(\lambda)$ is the right model

See R script



Optional references



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Estimating classification error rate: Repeated cross-validation, repeated hold-out and bootstrap.

Computational Statistics & Data Analysis, 53 (11): 3735-3745



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